X(4th Sm.)-Statistics-H/CC-9/CBCS

# 2022

## **STATISTICS** — HONOURS

### Paper : CC-9

# (Statistical Inference-I and Sampling Distribution)

#### Full Marks : 50

The figures in the margin indicate full marks. Candidates are required to give their answers in their own words as far as practicable.

Answer any five questions from question nos. 1-8. 2×5

- 1. Let  $X_1, ..., X_n$  be a random sample from  $N(\theta, 1)$  choose d so that test based on  $T_n = (\# \text{ of } |X_i| \le d)/n$  for testing  $H_0: \theta = 0$  against  $H_1: \theta = 10$  has both types of errors almost zero.
- 2. Let  $X_1, ..., X_n$  be a random sample from Bernoulli (p) population,  $P \in (0,1)$ . Define *p*-value of the test based on  $T_n = n \{\# \text{ of } X_i = 0\}$ , for testing  $H_0: p = P(X_1 = 0) = \frac{1}{2}$  against  $H_1: p > \frac{1}{2}$ .
- 3. Let  $X_1, ..., X_{n_1}, X_{n_{1-1}}, ..., X_n$  be IID  $N(\theta, \sigma^2), \theta \in \mathbb{R}, \sigma > 0$ . Consider statistics

$$T_{1n} = \frac{\overline{X}_n - \theta}{\left[\frac{1}{n-1}\sum_{i=1}^n (X_i - \overline{X}_n)^2\right]^{\frac{1}{2}}} \text{ and } T_{2n} = \frac{\frac{1}{n-n_1}\sum_{i=1}^n X_i - \theta}{\left[\frac{1}{n_1-1}\sum_{i=1}^n (X_i - \overline{X}_{n_1})^2\right]^{\frac{1}{2}}}, \overline{X}_{n_1} = \frac{1}{n_1}\sum_{i=1}^n X_i, \overline{X}_n = \frac{1}{n_1}\sum_{i=1}^n X_i$$

Are sampling distribution of  $T_{1n}$  and  $T_{2n}$  same? Justify with derivations.

- 4. Based on a random sample  $\{X_1, ..., X_{10}\}$  from N(0, 1) construct  $T(X_1, ..., X_{10})$  such that T follows a  $\chi^2$  distribution with 3 d.f. Use all the observations for constructing T.
- 5. Let  $X_1, ..., X_n$  be N(0, 1), show that  $E(X_{(1)} + X_{(n)}) = 0$ .
- 6. Let  $X_1, ..., X_n$  be a random sample from  $N(\theta, 1), \theta \in \mathbb{R}$ . Consider the confidence interval  $S = (\overline{X}_n d, \overline{X}_n + d)$  for  $\theta$  where  $d > \frac{Z_{\alpha/2}}{\sqrt{n}}$ ,  $Z_{\alpha/2}$  is the upper level  $\alpha/2$  point of N(0, 1) distribution

and  $\alpha \in (0, 1)$ . If observed interval  $(\overline{x}_n - d, \overline{x}_n + d)$  contains  $\theta$ , is it possible to reject alternative at level  $\alpha$  for testing  $H_0: \theta = 0$  against  $H_1: \theta \neq 0$  based on critical region S? Justify with necessary derivations.

#### **Please Turn Over**

# X(4th Sm.)-Statistics-H/CC-9/CBCS

- 7. Let  $\{X_1, ..., X_{20}\}$  be IID  $F_{2,3}$  distribution. Find the value of  $E\left[\sum_{i=1}^{5} X_i^{-1} / \sum_{i=1}^{20} X_i^{-1}\right]$ .
- 8. If  $X \sim t_n$  (t-distribution with n.d.f.), what is the distribution of  $\left(1 + \frac{X^2}{n}\right)^{-1}$ ? Derive in details.

Answer any two questions from question nos. 9-11.

- 9. Let  $X_1, ..., X_{n_1}$  and  $Y_1, ..., Y_{n_r}$  be two independent random samples from Poisson  $(\lambda_1)$  and Poisson  $(\lambda_2)$ respectively. Perform a test based on *P*-value for testing  $H_0: \lambda_1 = \lambda_2$  against  $H_1: \lambda_1 > \lambda_2$ .
- 10. Let  $X_1, ..., X_n$  be independent and identically distributed continuous random variables with distribution function F(x). Then show that

 $\lim_{n \to \infty} P_F(n(F(X_{(n)}) - F(X_{(n-1)})) \le t) = e^{-1} \forall t > 0, \text{ when } X_{(i)} \text{ is the } i\text{-th order statistics.}$ 

- 11. Let X and Y be independent variables having common distribution Exponential ( $\lambda$ ),  $\lambda > 0$ . Find (a) conditional distribution of w given u and hence
  - (b) marginal distribution of w, where u = X + Y and w = X Y.

Answer any three questions from question nos. 12-16.

- 12. (a) If random variables  $X_1$  and  $X_2$  are independent and each follows  $X^2$ -distribution with *n* d.f., show that  $T = \frac{\sqrt{n}(X_1 - X_2)}{2\sqrt{X_1X_2}}$  follows a student's *t*-distribution with n - 1 d.f. and distribution of T is independent of  $X_1 + X_2$ .
  - (b) Let  $X_1, ..., X_{n_1}$  and  $Y_1, Y_2, ..., Y_{n_2}$  be independent random samples. from  $N(\mu_1, \sigma_1^2)$  and  $N(\mu_2, \sigma_2^2)$ respectively and all the parameters are unknown. Find a confidence interval of  $\frac{\sigma_1^2}{\sigma_2^2}$  with confidence coefficient  $(1 - \alpha)$ .  $\alpha \in (0, 1)$ . 6+4
- 13. (a) Derive the m.g.f. of  $\chi^2$  distribution with *n* d.f. Hence or otherwise show that  $\mu_{r+1} = 2r(\mu_r + \mu_{r-1}), r \ge 1$ , where  $\mu_k = E(\chi_n^2)^k, k \ge 1$ .

(2)

5×2

### X(4th Sm.)-Statistics-H/CC-9/CBCS

(b) Let  $(Z_1, Z_2)$  have bivariate normal distribution with  $E(Z_1) = E(Z_2) = 0$ ,  $V(Z_1) = V(Z_2) = 1$  and correlation  $(Z_1, Z_2) = \rho$ . Suppose  $Z_1$  and  $Z_2$  are unobservable and the observable random variables are

$$X_{i} = \begin{cases} 0, & \text{if } Z_{i} \le 0\\ 1, & \text{if } Z_{i} > 0 \end{cases} \quad i = 1, Z$$

Let  $\tau$  be the correlation coefficient between  $X_1$  and  $X_2$ . Prove that  $\rho = \sin(\pi \tau/2)$ . 5+5

- 14. (a) For a bivariate sample {(Y<sub>i</sub>, X<sub>i</sub>); i = 1,...,n}, consider regression model Y<sub>i</sub> = βX<sub>i</sub> + ε<sub>i</sub>, i = 1,...,n, where ε<sub>i</sub> is independent of x<sub>i</sub> and ε<sub>i</sub>'s are independent and identically distributed as N(0, σ<sup>2</sup>). Derive a test for H<sub>0</sub>: β = 0 against H<sub>1</sub>: β ≠ 0.
  - (b) If  $X_1, ..., X_n$  be a random sample from  $N(\mu, \sigma^2)$ , find the sampling distribution of

$$\sqrt{\frac{n}{n-1}} \left(\overline{X}_n - X_n\right) / \sqrt{\left\{ (n-1)S_n^2 - \frac{n}{n-1}(X_n - \overline{X}_n)^2 \right\} / (n-2)},$$
  
where  $\overline{X}_n = \frac{1}{n} \sum_{1}^n X_i$  and  $S_n^2 = \frac{1}{n-1} \sum_{1}^n \left(X_i - \overline{X}_n\right)^2.$  6+4

- 15. (a) Let the random variables X and Y be distributed as  $\chi^2_{n_1}$  and  $F_{n_1,n_2}$  respectively. For any  $\alpha \in (0, 1)$ ,  $\chi^2_{\alpha, n_1}$  and  $F_{\alpha, n_1, n_2}$  be defined by  $P(X \ge \chi^2_{\alpha, n_1}) = p(Y \ge F_{\alpha, n_1}, n_2) = \alpha$ . Then show that for large  $n_2$ ,  $\chi^2_{\alpha, n_1} \approx n_1 F_{\alpha, n_1, n_2}$ .
  - (b) For a bivariate sample  $\{(X_{1i}, X_{2i}), i = 1, ..., n\}$  from bivariate normal distribution with unknown parameters  $(\mu_1, \mu_2, \sigma_1^2, \sigma_2^2, \rho)$ , derive a test for  $H_0: \sigma_1^2 = \sigma_2^2$  against  $H_1: \sigma_1^2 \neq \sigma_2^2$ . 5+5
- 16. (a) Let  $X_1, ..., X_n$  be i.i.d. random variables with continuous d.f. F and let  $X_{(i)} < ... < X_{(n)}$  be the order statistics. If  $M_0$  be the unique population median, then show that

$$P_{F}(X_{(r)} \le M_{0} \le X_{(s)}) = \left| \sum_{k=r}^{s-1} \binom{n}{k} \right| \left(\frac{1}{2}\right)^{n}, r < s.$$

Hence or otherwise find a confidence interval of  $M_0$  with coverage probability at least  $1 - \alpha$  for some  $\alpha \in (0, 1)$ .

(b) Find the mean and variance of Student's *t*-distribution. Show that its density tends to N(0, 1) as degrees of freedom becomes large. 6+4

(3)